

# Distributed Hypothesis Testing

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We model a network composed of  $n$  agents as a graph  $G = \{V, E\}$ .  $V = \{1, 2, \dots, n\}$  is the set of vertices representing the agents.  $E \subseteq V \times V$  is the set of edges.  $(i, j) \in E$  if and only if sensor  $i$  and  $j$  can communicate directly with each other. We will always assume that  $G$  is undirected, i.e.  $(i, j) \in E$  if and only if  $(j, i) \in E$ . We further assume that there is no self loop, i.e.,  $(i, i) \notin E$ .

At each time, each sensor make an i.i.d. measurement  $y_i(k)$ . Consider the following two hypothesis:

$$H_0 : y_i(k) \sim \mathcal{N}(0, 1).$$

$$H_1 : y_i(k) \sim \mathcal{N}(1, 1).$$

We assume that each hypothesis is true with 0.5 probability.

## 1 Centralized Detector

The optimal centralized detector is a Naive Bayes detector. Define the average to be

$$\alpha(k) = \frac{1}{n(k+1)} \sum_{t=0}^k \sum_{i=1}^n y_i(k).$$

Hence, the centralized detector is

$$f(\alpha(k)) = \begin{cases} 0 & \text{if } \alpha(k) \leq 0.5 \\ 1 & \text{if } \alpha(k) > 0.5 \end{cases}$$

Define the probability of error of such detector to be  $P_c(k)$ , then

$$P_c(k) = 0.5P(\alpha(k) \leq 0.5|H_1) + 0.5P(\alpha(k) > 0.5|H_0) = P(\alpha(k) > 0.5|H_0)$$

We use large deviation theory to characterize  $P_c(k)$ . Suppose  $x_i(k) \sim \mathcal{N}(0, 1)$ , then the moment generating function is given by

$$M(\theta) = \int_{-\infty}^{\infty} \exp(\theta t) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{t^2}{2}\right) dt = \exp\left(\frac{\theta^2}{2}\right) \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{(t-\theta)^2}{2}\right) dt = \exp\left(\frac{\theta^2}{2}\right).$$

Hence, the log-moment generating function is

$$\Lambda(\theta) = \log M(\theta) = \frac{\theta^2}{2},$$

and

$$I(0.5) = \sup_{\theta} 0.5\theta - \Lambda(\theta) = \frac{1}{8}.$$

Hence, we know that  $P_c(k) \sim \exp(-nk/8)$ .

## 2 Distributed Detection

Let  $A$  be a consensus matrix that is compatible with the topology  $G$ , such that

- $A$  has an eigenvalue of 1 and all the other eigenvalues of  $A$  are strictly inside the unit disk.
- $\mathbf{1}$  is both a left and right eigenvector of  $A$ .

Define  $J = \mathbf{1}\mathbf{1}^T/n$ , then  $A^k \rightarrow J$  as  $k \rightarrow \infty$ .

Define the state of sensor  $i$  at time  $k$  to be  $x_i(k)$ . The sensor update equation can be written as

$$\begin{aligned} x_i(k)^+ &= \frac{k}{k+1}x_i(k) + \frac{1}{k+1}y_i(k), \\ x_i(k+1) &= a_{ii}x_i(k)^+ + \sum_{j \in N_i} a_{ij}x_j(k)^+. \end{aligned}$$

Hence,

$$\begin{aligned} x(k)^+ &= \frac{k}{k+1}x(k) + \frac{1}{k+1}y(k), \\ x(k+1) &= Ax(k)^+. \end{aligned}$$

Let us define

$$\bar{x}(k+1) = Jx(k+1) = \alpha(k)\mathbf{1}.$$

For each sensor, it implements a detector  $f_i$ , which is defined as

$$f_i(x_i(k)) = \begin{cases} 0 & \text{if } x_i(k) \leq 0.5 \\ 1 & \text{if } x_i(k) > 0.5 \end{cases}$$

Denote the probability of error for each individual detector as  $P_i(k)$ . By symmetry, we know that

$$P_i(k) = P(x_i(k) \geq 0.5 | H_0).$$

Clearly  $x_i(k) < 0.5$  if  $\alpha(k-1) < 0.5 - \delta$  and  $x_i(k) - \alpha(k-1) < \delta$ , for any  $\delta > 0$ . Hence,

$$P_i(k) = P(\alpha(k-1) \geq 0.5 - \delta | H_0) + P(x_i(k) - \alpha(k-1) \geq \delta | H_0).$$

For the first probability, we know that

$$I(0.5 - \delta) = \sup_{\theta} (0.5 - \delta)\theta - 0.5\theta^2 = 0.125 + \varepsilon,$$

where  $\varepsilon \rightarrow 0$  when  $\delta \rightarrow 0$ . Hence,  $P(\alpha(k-1) \geq 0.5 - \delta | H_0) \sim \exp(-(0.125 + \varepsilon)nk)$ .

Now let us look at  $k(x_i(k) - \alpha(k-1))$ . We know that

$$k(x(k) - \bar{x}(k)) = [(A - J)y(k-1) + (A - J)^2 y(k-2) + \dots + (A - J)^k y(0)].$$

Hence, under hypothesis  $H_0$ ,  $k(x(k) - \bar{x}(k))$  is Gaussian distributed with zero mean and with covariance:

$$(A - J)(A - J)^T + \dots + (A - J)^{k+1} ((A - J)^{k+1})^T \leq M.$$

Therefore,  $k(x_i(k) - \alpha(k-1))$  is Gaussian distributed with zero mean and a bounded variance.

$$P(x_i(k) - \alpha(k-1) \geq \delta | H_0) = \frac{1}{\sqrt{2\pi}} \int_{k\delta/\sigma(k)}^{\infty} \exp\left(-\frac{t^2}{2}\right) dt,$$

where  $\sigma(k)$  is the standard deviation of  $k(x_i(k) - \alpha(k-1))$ .

For any  $x > 0$ , we have that

$$\frac{1}{\sqrt{2\pi}} \int_x^{\infty} \exp\left(-\frac{t^2}{2}\right) dt \leq \frac{1}{\sqrt{2\pi}} \int_x^{\infty} \frac{t}{x} \exp\left(-\frac{t^2}{2}\right) dt \leq \frac{1}{x\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right).$$

Therefore, for large enough  $k$

$$P(x_i(k) - \alpha(k-1) \geq \delta | H_0) \leq \exp\left(-\frac{1}{2M_{ii}} k^2 \delta^2\right).$$

Hence, one can prove that for any  $\varepsilon > 0$ , and large enough  $k$ ,

$$P_i(k) \leq \exp(-(0.125 + \varepsilon)nk).$$

On the other hand,  $P_i(k) \geq P_c(k)$ , which implies that

$$P_i(k) \geq \exp(-(0.125 - \varepsilon)nk).$$

Hence,  $n/8$  is the rate function for  $P_i(k)$ .